

Estimating Parameters of the Norris-Landzberg Model under Gamma Distribution

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Abstract: With the increase quality of products, the traditional life-testing method is too time-consuming. Therefore, the accelerated life-testing (ALT) is widely used. If the product operating conditions continuously, fluctuate the cvclic accelerated life-testing (CALT) is the considered. Among CALTs. Norris-Landzberg model is commonly used in modeling the growing crack resulting from the thermal cyclic stress. In the parameter estimate of this model, many researchers employed have linear regression models without a proper parametric assumption, which may impair its statistical integrity. Therefore, this paper propose a Gamma distribution assumption, which will ensure the unbiasedness and consistency of estimator. The corresponding inference is also shown.

Keywords: Norris-Landzberg Model; Linear Regression; Statistical Inference.

1.Introduction

In today's fiercely competitive global market, products are manufactured to meet stringent quality standards, resulting in extended product lifespans. Traditional life-testing methods, which rely on observing failures under standard operating conditions, can be impractical due to the lengthy period required to collect meaningful data. To address this challenge, accelerated life-testing (ALT) techniques have become a popular approach to expedite the data collection process. Comprehensive discussions on ALT can be found in the literature [1-6].

While the majority of research has concentrated on constant stress ALT, the cyclic accelerated life-testing (CALT) method has been introduced for scenarios where product operating conditions fluctuate continuously. This approach has been

successfully applied across various industries, including the reliability assessment of engines [7] and electric drive systems [8]. CALT several factors incorporates such as temperature cycling, vibration, and humidity fluctuations. Notably, the Coffin-Manson model [9], which focuses on the temperature range, is commonly used for modeling temperature cycling. However, this model account for the maximum does not temperature and cycling frequency, which are also critical parameters. To provide a more comprehensive framework for CALT analysis, the Norris-Landzberg model [10] has been proposed. Extensive researches are based on the Norris-Landzberg model, see [11-13].

In the parameter estimation process of the Norris-Landzberg model, many researchers have employed linear regression models without a proper parametric assumption, leading to estimators that may lack statistical robustness. This could result in biased and inconsistent estimators. То ensure the statistical integrity of the least squares method, a Gamma distribution we introduce assumption in this paper. It is observed that parameter of the Gamma shape the distribution does not affect the inference of the Norris-Landzberg model, hence the shape parameter is assumed to be known.

The structure of this paper is as follows: Section 2 outlines the methodology, including the Norris-Landzberg model and the least squares method with Gamma distribution family. Section 3 details the simulation and analysis procedures, covering simulation settings, the accuracy of estimates, the asymptotic normality of estimates, and the evaluation of normal confidence intervals. Finally, Section 4 presents the concluding remarks.

2.Methodology

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2.1 Norris-Landzberg model

The Norris-Landzberg model, proposed by [10], is widely used in modeling the growing crack resulting from the thermal cyclic stress. It assumes that the number of cycle to failure, denoted by N, is influence by the temperature range ΔT , the highest temperature in the cycle T_{max} , and the cycling frequency f. The relationship is shown in the following formula

$$\mathbf{N} = A f^{-a} (\Delta T)^{-b} G(T_{max}) \tag{1}$$

where A > 0 represents the constant depends on the material,

a represents the exponent of cycling frequency, typically near $-\frac{1}{3}$,

b represents the exponent of the temperature range, which depends on the test environment, $G(T_{max}) = exp\{\frac{E_a}{K^*T_{max}}\}$ represents Arrhenius term with $T = T_{max}$, where $K^* = 8.623 \times$ $10^5 eV/K$ is the Boltzmann's constant and

$$T_{ij} = \frac{N_{ij}}{f_i} = A f_i^{-(a+1)} (\Delta T_i)^{-b} G(T_{i,max}), i = 1, 2, ..., n, j = 1, 2, ..., N_i$$
(4)

$$i = 1, 2, ..., n, j = 1, 2, ..., N_i$$
a Gamma distribution family $Gamma(k, \beta_i)$

where $\{T_{ij}|i=1,2,\ldots,n, j=1,2,\ldots,N_i\}$ represent the *i*th test environment's failure time. The *i*th stress factor, denoted by $S_i =$ $[f_i, \Delta T_i, T_{i,max}]$, represents the strength of stress in the *i*th test environment. The stresses in the test environments influence the lifetimes of test units, which motive us to use

satisfies the conditions of Gauss-Markov

Theorem. Hence, taking $\theta = [\ln(A), (a +$

1), b, E_a] as the parameter vector, we can

 $\boldsymbol{Q}(\boldsymbol{\theta}) = \sum_{i=1}^{n} \sum_{j=1}^{N_i} \left(\ln(t_{ij}) - \widehat{\ln(t_{ij})} \right)^2 \quad (6)$

where $\ln(t_{ij}) = \ln(A) - (a+1)\ln(f_i) - b\ln(\Delta T_i) + E_a \frac{1}{K^* T_{i,max}}$ is the estimated value

of $\ln(t_{ii})$. By the least squares method, we can minimize $Q(\boldsymbol{\theta})$ the by taking $\boldsymbol{\theta} = \widehat{\boldsymbol{\theta}}$.

derive the loss function

$$\ln(t_{ij}) = \ln(A) - (a+1)\ln(f_i) - b\ln(\Delta T_i) + E_a \frac{1}{K^* T_{i,max}} + \epsilon_{ij}, i = 1, 2, \dots, n, j = 1, 2, \dots, N_i$$
(5)

3. Simulation and Analysis

log-linear model

1. This section discusses the simulation and analysis procedure, which is divided into 4 subsections:

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Coffin-Manson model in practice.

 $N = A(\Delta T)^{-b}$

Coffin-Manson-Arrhenius model

experiment.

Coffin-Manson-Arrhenius

Coffin-Manson model

 $0, b \ge 0$ and $E_a \ge 0$.

Distribution Family

and Intelligent Technology (SDIT2024)

 E_a is the activation energy determined by

We also use the reduce model like

 $N = A(\Delta T)^{-b} G(T_{max})$

exclude the cycling frequency term since

a = 0. It can be further reduced to the

by taking $E_a = 0$. The subsequent discussion

will focus on the realistic case with $a \leq$

2.2 Least Square Method with Gamma

It is assumed that the real failure time T is

visible for all test units. Then the formulation of the Norris-Landzberg model is changed by

to formulate the lifetime distributions. Here k

is assume to be known and β_i is assumed to

be relevant to the *i* th test environment's

stresses. Since the shape parameter is not

influence by the environment's stresses, the

and

The

(2)

(3)

(4)

model

- 2. The simulation settings.
- 3. The accuracy of estimates.
- 4. The asymptotic normality of estimates

5. The evaluation of normal confidence intervals

3.1 The Simulation Settings

Setting	Norris–Landzberg model								
	i=1	i=2	i=3	i=4	i=5	i=6	i=7	i=8	
f	72	48	72	48	72	48	72	48	
ΔT	120K	100K	100K	120K	120K	100K	100K	120K	
T _{max}	393K	373K	373K	393K	373K	393K	393K	373K	
Parameter setting									
$E_a = 0.17E\nu, \ A = 10^4 = exp(9.2103), \ a = -\frac{1}{3}, \ b = 1.9$									
Setting		Coffin–Manson–Arrhenius model							
	i=1	i=2	i=3	i=4	i=5	i=6	i=7	i=8	

Table 1. Settings Used in the Simulation Study

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f	72	48	72	48	72	48	72	48			
ΔT	120K	100K	100K	120K	120K	100K	100K	120K			
T _{max}	393K	373K	373K	393K	373K	393K	393K	373K			
Parameter setting											
$E_a = 0.17Ev, \ A = 10^4 = exp(9.2103), \ a = 0, \ b = 1.9$											
Satting		Coffin–Manson model									
Setting	i=1	i=2	i=3	i=4	i=5	i=6	i=7	i=8			
f	72	48	72	48	72	48	72	48			
ΔT	120K	100K	100K	120K	120K	100K	100K	120K			
T _{max}	<u>393K 373K 373K 393K 373K 393K 393K 373K</u>										
Parameter setting											
$E_a = 0Ev, \ A = 10^4 = exp(9.2103), \ a = 0, \ b = 1.9$											
Setting	Normal Condition										
$f = 1, \Delta T = 30K, T_{max} = 303K$											

We consider the Norris-Landzberg model along with the degraded cases of the Coffin-Manson-Arrhenius model and the Coffin-Manson model. The detailed settings are shown in the Table 1. The Gamma distribution $Gamma(6, \beta_i)$ is utilized in the

$$\beta_i = A \exp(-(a+1)\ln(f_i) - b\ln(\Delta T_i) + E_a \frac{1}{\mu^* \tau} - \psi(k))$$
(7)

where $\psi(k)$ represents the digamma function evaluated at k.

3.2 The Accuracy of Estimates	
Table 2. Bias and MSE of the est	timators

$Gamma(6, \beta_i)$		Norris-Landzberg model							
		Α	а	b	E _a				
	Bias	-0.0303	-0.0034	-0.0011	0.0003				
$N_{i} = 20$	MSE	5.1401	0.0274	0.1345	0.0018				
	Bias	0.0227	-0.0038	0.0009	0.0001				
$N_{i} = 40$	MSE	2.6920	0.0132	0.0725	0.0009				
	Bias	-0.0049	-0.0028	0.0027	0.0002				
$N_{i} = 80$	MSE	1.2432	0.0066	0.0342	0.0005				
Gamma	$(6, \beta_i)$	Coffin-Manson-Arrhrnius model							
		Α	а	b	E_a				
	Bias	0.0753	/	0.0147	-0.0002				
$N_{i} = 20$	MSE	4.7756	/	0.1398	0.0018				
	Bias	-0.0386	/	-0.0110	0.0007				
$N_i = 40$	MSE	2.3592	/	0.0673	0.0009				
	Bias	0.0058	/	-0.0015	-0.0004				
$N_{i} = 80$	MSE	1.1960	/	0.0346	0.0004				
Gamma	(6, β _i)	Coffin-Manson model							
		Α	а	b	E _a				
	Bias	-0.0221	/	-0.0048	/				
$N_{i} = 20$	MSE	3.1532	/	0.1429	/				
	Bias	-0.0607	/	-0.0032	/				
$N_i = 40$	MSE	1.4714	/	0.0667	/				
	Bias	-0.0160	/	-0.0034	/				
$N_i = 80$	MSE	0.7774	/	0.0353	/				

In this section, the mean square error (MSE) is used to evaluate the performance of $\hat{\theta}$, defined as:

 $MSE(\hat{\theta}) = Bias^2(\hat{\theta}) + Var(\hat{\theta}).$ (8) The simulation was performed 10,000 times

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simulation, while k=6 is the fixed shape parameter and β_i represents the scale parameter in ith test group. For the Gamma distribution, the scale parameters β_i for i = $1, 2, \ldots, n$ are expressed as

$$A\exp(-(a+1)\ln(f_i) - b\ln(\Delta T_i) + E_a \frac{1}{K^* T_{imag}} - \psi(k)) \tag{7}$$

with $N_i = 20,40,80$ respectively and the findings were meticulously analyzed and detailed in Table 2. As the quantity of units increases, both the bias and the MSE for each model show a decline, signifying that the estimator is both unbiased and consistent.

3.3 The asymptotic normality of estimates

If the condition of Gauss-Markov Theorem is satisfied, the estimators derived by the least squares method have the asymptotic normality. When the lifetimes follow the Gamma family, the G-M Theorem is satisfied, leading to the asymptotic normality of estimators. It can also be verified by the QQ-plots in the Figure 1. It shows the QQ-plots of each estimator when $N_i = 20,40,80$. It can be seen that even in the case that $N_i = 20$, the normality of estimators still maintain, which is a basis of the normal confidence interval of parameters.

3.4 The evaluation of the normal confidence intervals

Table 3. Coverage Probabilities	of	the
Normal CIs		

Gamma(6, β_i)		Norris-Landzberg model							
		Α		а		b		E_a	
$N_i = 20$	90%	0.8	8945		0.9001		0.9023		0.9045
	95%	0.9	9427		0.9456		0.9502		0.9503
$N_i = 40$	90%	0.8	8926		0.8977		0.8976		0.8952
	95%	0.9	9562		0.9512		0.9504		0.9521





4. Conclusions

This paper propose a Gamma distribution family assumption, and under this assumption

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When the normality of estimators is satisfied, we can use normal confidence interval to make interval estimate by formula

$$[\underline{\theta_k}, \overline{\theta_k}] = [\widehat{\theta_k} \pm z_{1-\frac{\alpha}{2}} \sqrt{[(A^T A)^{-1}]_{k,k}}]$$
(9)

where A represent coefficient matrix. The simulation was performed 10,000 times with $N_i = 20,40,80$ respectively and we calculate the coverage probability of the 90% and 95% confident interval. The detailed results are shown in the Table 3. It can be seen that the coverage probabilities in every groups are consistent with the nominal levels, that means, the normal interval is effective in the confidence interval construction.

Fa

0 1 2

0

0

2 3

1

2



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method will be effective. In addition, this paper also explain the asymptotic normality of estimators, which ensures the effectiveness of the normal confidence interval. In reality, there are many other hypotheses with the same nature, which will become our future topic.

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